Advanced computational statistics, lecture 6

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Course schedule

- Topic 1: **Gradient based optimisation**
- Topic 2: **Stochastic gradient based optimisation**
- Topic 3: **Gradient free optimisation**
- Topic 4: **Optimisation with constraints**
- Topic 5: **EM algorithm and bootstrap**
- Topic 6: **Simulation of random variables**
- Topic 7: **Numerical and Monte Carlo integration; importance sampling**

Course homepage: http://www.adoptdesign.de/frankmillereu/adcompstat2023.html

Includes schedule, reading material, lecture notes, assignments

Simulation in Statistics

- Computer-generated random variables
- Purpose:
  - Simulate a situation where a statistical model can be assumed
  - Simulate situation repeatedly to investigate properties of estimators, confidence intervals, significance tests
    - Example: power of a test in situations where assumptions are not fulfilled
  - Perform Monte Carlo integration

- Problem: Given a density f of a target distribution, generate random draws $X_1, \ldots, X_n$ which follow the target distribution

Random variables from familiar distributions

- Computer-generated random variables are not really random but deterministic (Gentle, Härdle, Mori, 2012, Ch.3)
- Algorithms are used such that the deterministic nature is not visible, and variables seem random

- Deterministic algorithm generates values between 0 and 1 which follow well independent draws from Unif[0,1]
- Then, random variables following other familiar distributions can be generated from Unif[0,1] and are implemented in statistical software, see Givens and Hoeting (2013), Tab. 6.1

Random variables of familiar distributions in R

- In R, random variables can be generated for a number of distributions, e.g:
- **rbeta, rcauchy, rchisq, rexp, rf, rgamma, rlnorm, rnorm, rt, runif, rweibull**
- **rbinom, rgeom, rhyper, rmultinom, rnbinom, rpois**

```
x <- rnorm(6, mean = 1.2, sd = 2)
x
[1]  3.8839870  2.8328797  3.5344539 -2.5464309  3.2059822  0.1872261

rbinom(25, size = 3, prob = 0.25)
[1] 1 2 0 0 0 0 0 2 3 0 0 2 1 1 0 0 1 0 1 1 2 2 1 0 0
```

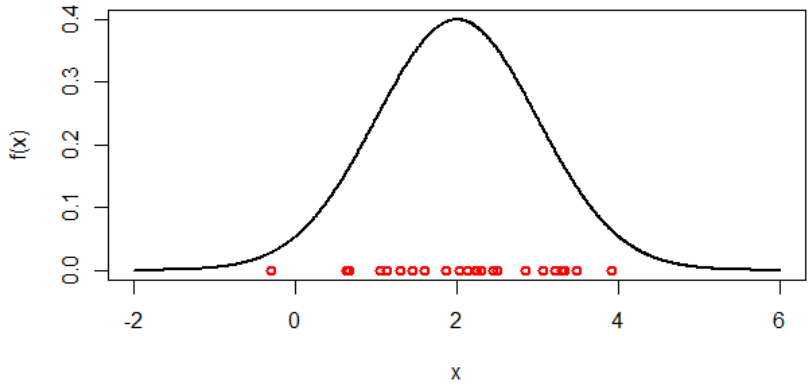
Random variables from non-familiar distributions

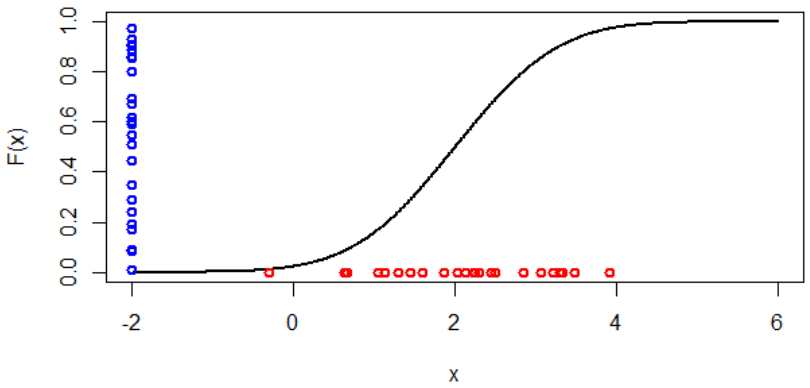
- Problem: Given a density *f* of a target distribution, generate random draws $X_1, \ldots, X_n$ which follow the target distribution
- Now: Density *f* of arbitrary form

- Methods we will consider:
  - Inverse transformation method
  - Rejection sampling
  - Composition sampling
  - Sampling importance resampling (SIR)
  - Markov chain Monte Carlo (MCMC)

Inverse transformation method

- Continuous random variable X with density f and distribution function F
- Then: $F(X)$ is uniformly distributed on [0,1]





- Therefore: if we can generate uniformly distributed random variables U, we can compute $X = F^{-1}(U)$ and obtain the desired sample

Inverse transformation method

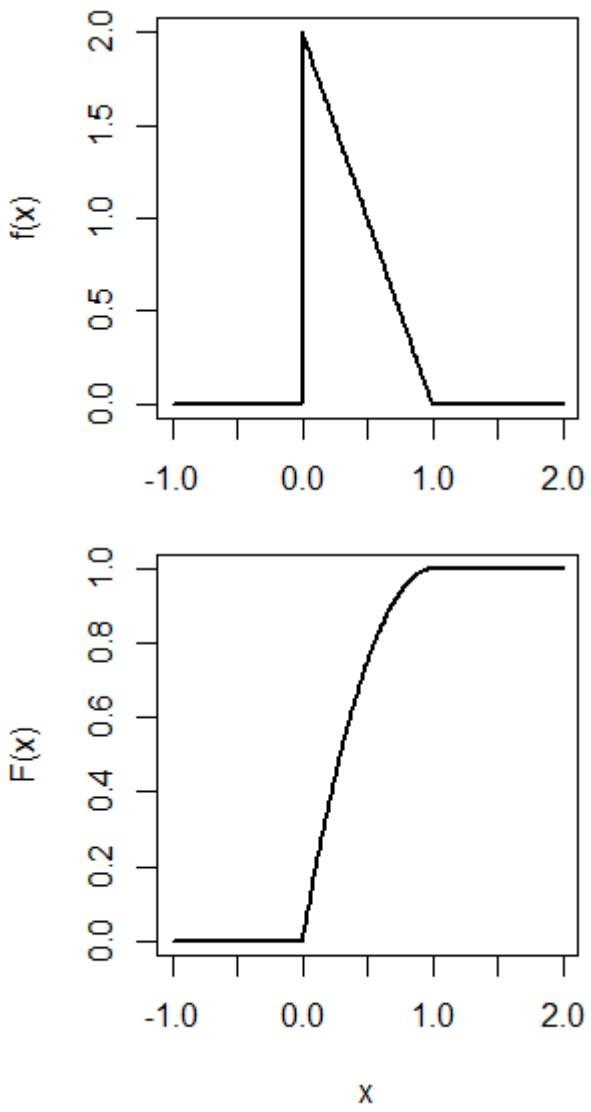
- Example 1: We want to generate random variables X with triangle distribution having density

$$f(x) = \begin{cases} 2 - 2x, & \text{if } 0 \le x \le 1, \\ 0, & \text{otherwise} \end{cases}$$

- We compute the distribution function:

$$F(x) = \int_{-\infty}^{x} f(t)\,dt = \begin{cases} 0, & \text{if } x < 0, \\ 2x - x^2, & \text{if } 0 \le x \le 1, \\ 1, & \text{if } x > 1. \end{cases}$$

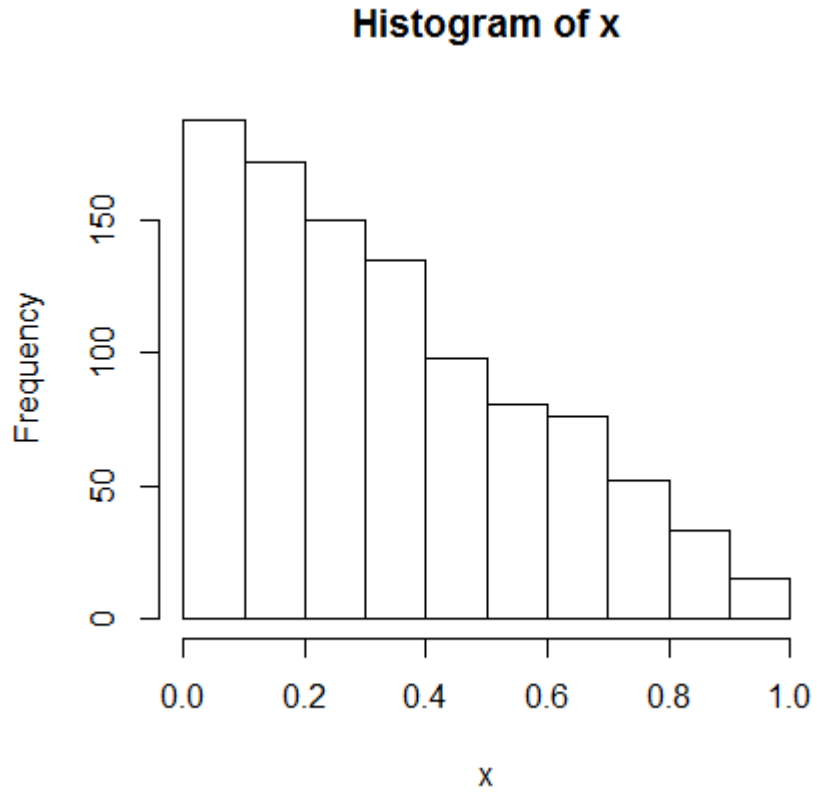
- The inverse distribution function is
  $F^{-1}(y) = 1 - \sqrt{1 - y}$
  since $y = 2x - x^2 \;\Leftrightarrow\; x^2 - 2x + y = 0 \;\Leftrightarrow$
  $x_{1,2} = 1 \pm \sqrt{1 - y} \;\Rightarrow 1 - \sqrt{1 - y}$



Inverse transformation method

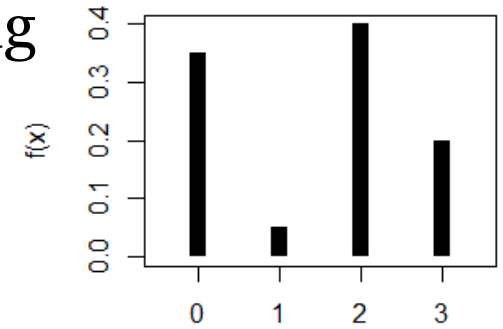
- 1000 random numbers for the triangle distribution can be generated by:

```
u <- runif(1000)
x <- 1-sqrt(1-u)
hist(x)
```

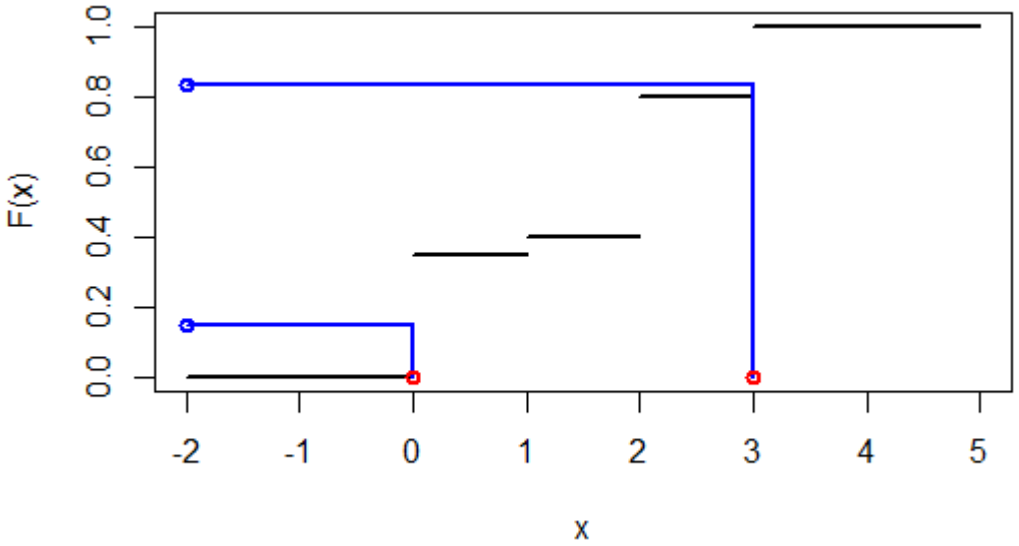


Inverse transformation – discrete random variables

- Example 2: We want to generate a random variable X being
  0 with probability 0.35,
  1 with probability 0.05,
  2 with probability 0.4,
  3 with probability 0.2

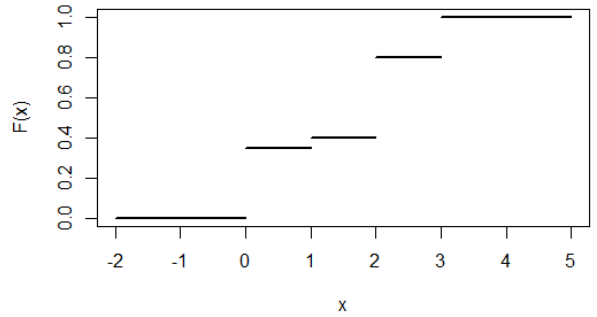


- F(x)=P(X≤x); how to apply inverse transformation method?



Inverse transformation – discrete random variables

- Example 2: We want to generate a random variable X being
  0 with probability 0.35,
  1 with probability 0.05,
  2 with probability 0.4,
  3 with probability 0.2
- How to apply inverse transformation method?
- Generate U~Unif[0,1]
- If $U \leq 0.35$, then $X = 0$,
  if $0.35 < U \leq 0.4$, then $X=1$,
  if $0.4 < U \leq 0.8$, then $X=2$,
  if $0.8 < U$, then $X=3$.




```
u <- runif(100000)
x <- (u>0.35)+(u>0.4)+(u>0.8)
```

This is 1 if the condition in (...) is true, otherwise it is 0

Inverse transformation method

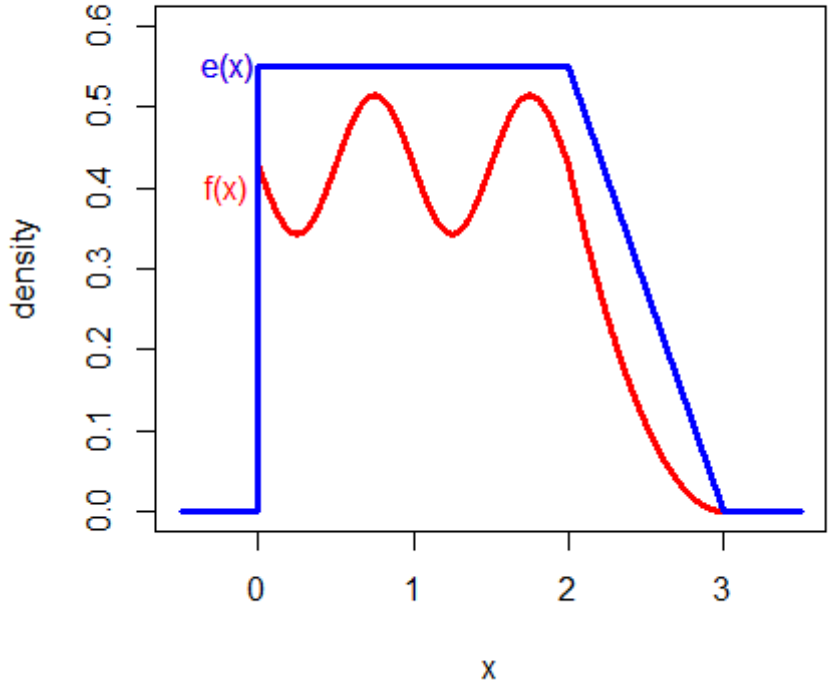
- Inverse transformation worked well in preceding examples
- In general, drawbacks are:
  - Computation of F^{-1} might be difficult
  - Not possible to generalize to multiple dimensions
  - Often less efficient as alternatives

Rejection sampling

- Problem: Given a density f of a target distribution, generate random draws $X_1, \ldots, X_n$ which follow the target distribution
- It can be difficult to sample with respect to f
- Situation: There is another density g which can be sampled from and which is after scaling larger than f for all x,

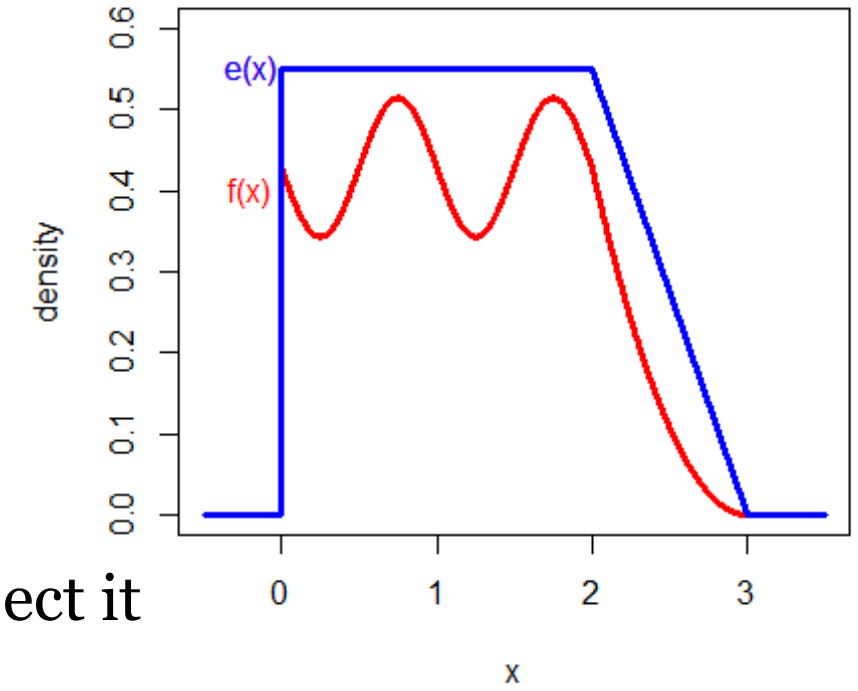
  $$e(x) = g(x)/\alpha \geq f(x)$$

  for all x and some $\alpha < 1$
- $e(x)$ is called ”envelope”



Rejection sampling

- $e(x)=g(x)/\alpha \geq f(x)$ for all x and some $\alpha < 1$



- Rejection sampling algorithm:

1. Sample Y~g
2. Sample U~Unif(0,1)
3. If U≤f(Y)/e(Y), accept Y; set X=Y; otherwise reject it
4. If more samples desired go to 1.

Example (for picture above): Y=2.21; f(Y)=0.267, e(Y)=0.435, f(Y)/e(Y)=0.616; sample U; If U≤0.616, use Y, otherwise reject it

Rejection sampling

1. Sample Y~g=eα
2. Sample U~Unif(0,1)
3. If U≤f(Y)/e(Y), accept Y; set X=Y; otherwise reject it
4. If more samples desired, go to 1

Example (for picture above):

(Y_1,U_1)=(2.21,0.492) ➔ U_1<0.616 ➔ accept Y_1

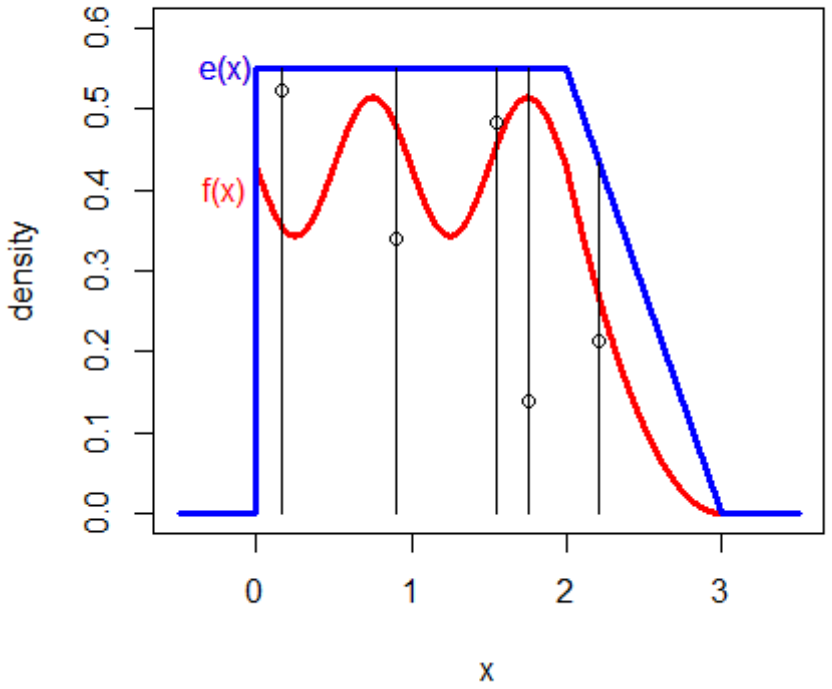
(Y_2,U_2)=(0.17,0.952) ➔ U_2>f(0.17)/e(0.17) ➔ reject Y_2

(Y_3,U_3)=(1.76,0.250) ➔ U_3<f(1.76)/e(1.76) ➔ accept Y_3

(Y_4,U_4)=(1.55,0.880) ➔ U_4>f(1.55)/e(1.55) ➔ reject Y_4

(Y_5,U_5)=(0.90,0.619) ➔ U_5<f(0.90)/e(0.90) ➔ accept Y_5

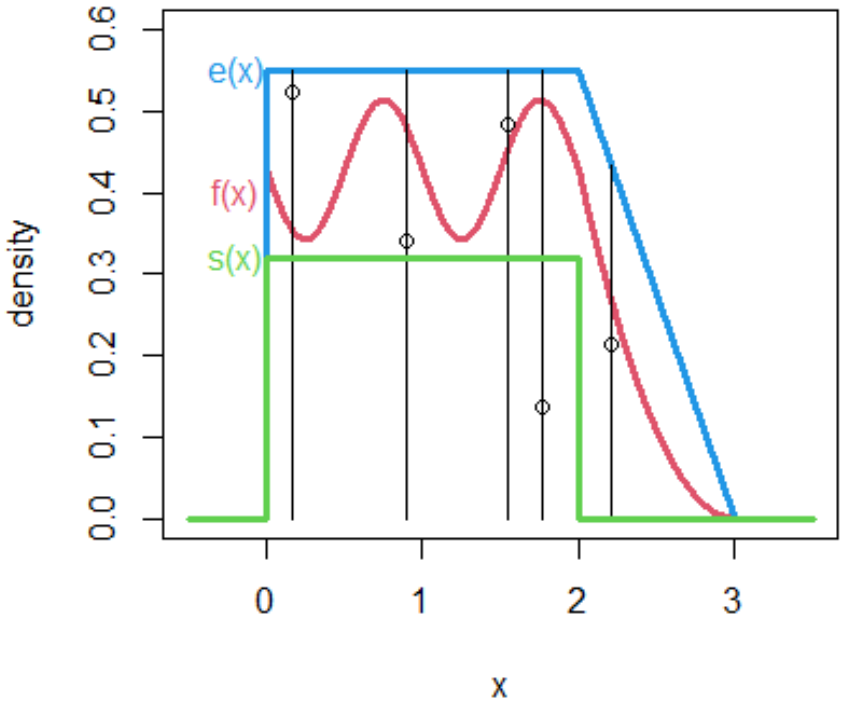
➔ use (2.21, 1.76, 0.90)



Squeezed rejection sampling

- $e(x)=g(x)/\alpha \geq f(x)$ for all x and some $\alpha < 1$
- **Squeezing function $s(x)$, $s(x) \leq f(x)$**
- **Squeezed** rejection sampling algorithm:

1. Sample Y~g
2. Sample U~Unif(0,1)
3. **If U≤s(Y)/e(Y), accept Y; set X=Y; go to 5**
4. If U≤f(Y)/e(Y), accept Y; set X=Y
5. If more samples desired go to 1.



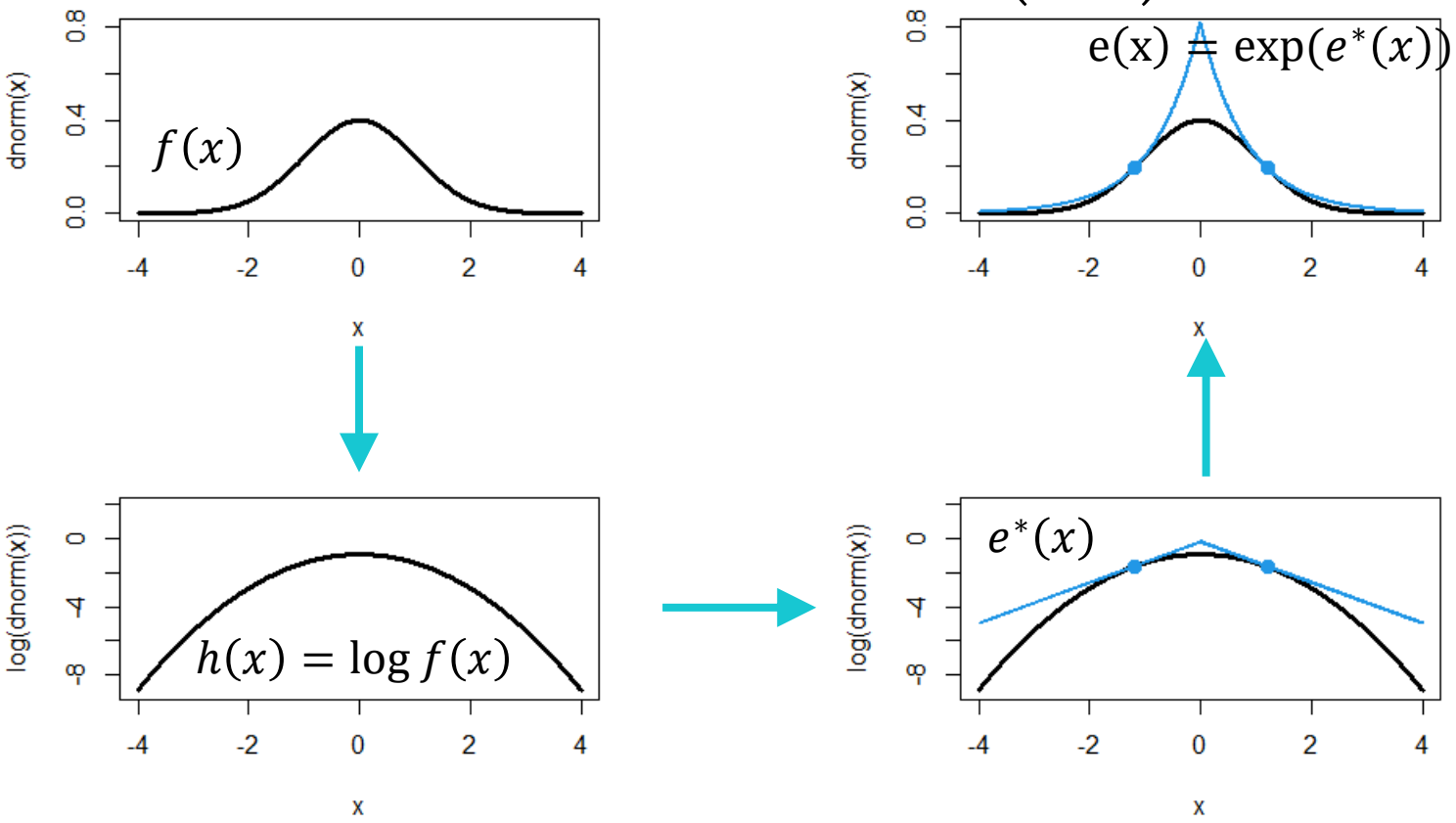
Example (for picture above): Y=0.90; s(Y)=0.32, e(Y)=0.55, s(Y)/e(Y)=0.582; sample U; If U<0.582, use Y, otherwise compute f(Y)=0.479, f(Y)/e(Y)=0.871, and use Y if U<0.871, otherwise reject

Adaptive (squeezed) rejection sampling

- An automated generation of envelope and squeezing functions can be good
- Adapt (improve) these functions where it is necessary
- Assumption: f log-concave, continuous, differentiable, $f>0$ on an interval I
- Start with grid $T_k = \{x_1, \ldots, x_k\}$ of points on I; consider $h=\log(f)$
- The tangents of the concave h in x_i form an upper hull $e^*(x)$ of h,
  ➔ $e(x) = \exp(e^*(x))$ is an envelope
- The interpolations between the x_i forms a lower hull $s^*(x)$ of h,
  ➔ $s(x) = \exp(s^*(x))$ is a squeezing function
- If x was rejected or (in the case of squeezing) if x is accepted in Step 4 $[s(x)/e(x)<U\leq f(x)/e(x)]$, then the point x is added to T_k ➔ T_{k+1}

Adaptive rejection sampling

- Example 3: N(0,1), $f(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$, $T_2 = \{x_1, x_2\}$, $x_2 = -x_1 = 1.2$



Proportion of waste:
$1\text{-}1/\int_{-\infty}^{\infty} e(x)\,dx$
Here: 1-1/1.366=0.268

Adaptive rejection sampling

- An adaptive rejection sampling version exists which does not require the derivative of h (secants instead of tangents are used, see Givens and Hoeting (2013; Chapter 6.2.3.2)

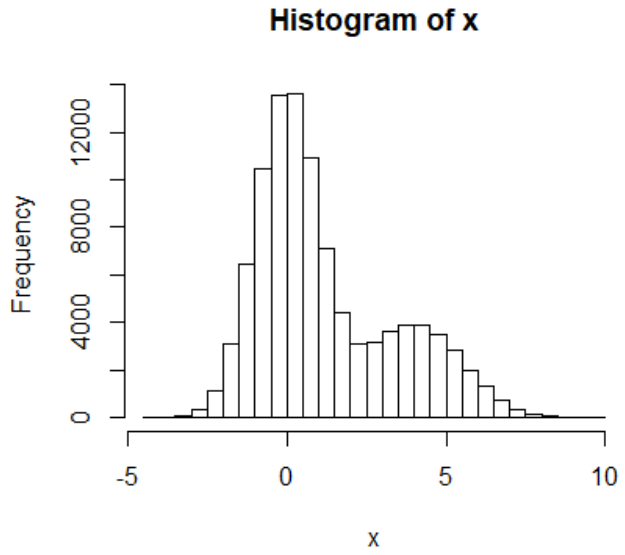
- Adaptive rejection sampling can be used for multidimensional cases, for example as subroutine in Gibbs sampling

- Many densities are log-concave, but some are not; non-log-concave densities can be handled by combining it with a Metropolis step

Composition sampling

- A finite mixture distribution can be generated by:
  - simulating the group-membership using the discrete distribution for mixing parameters
  - simulating the distribution of this group's distribution
- See Gentle, Härdle, Mori (2012), Section 3.8.7
- Ex. 4: X normal mixture of N(0,1) and N(4,1.5^2) with mixing parameter 0.7 and 0.3, respectively

```
g <- rbinom(100000, size = 1, prob = 0.3)
x <- rnorm(100000, mean = 4*g, sd = 1+0.5*g)
hist(x, breaks = 25)
```



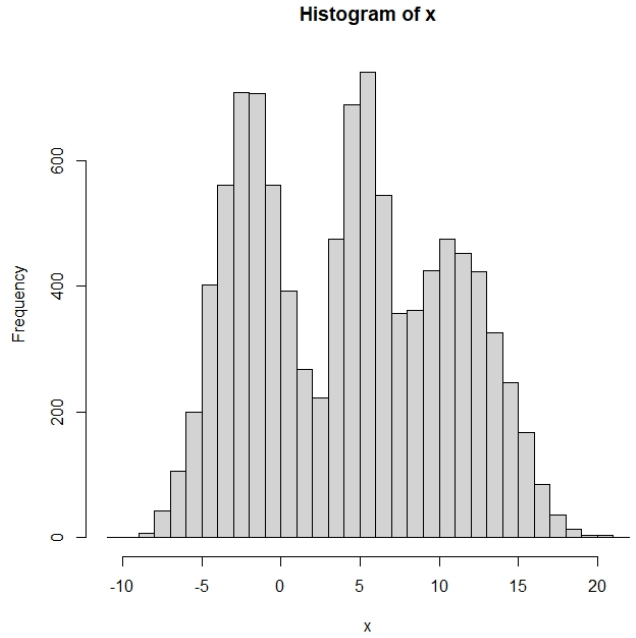
Composition sampling

- More flexible code for simulating a finite mixture distribution (e.g., a finite normal mixture) with composition sampling:
  - Define mean, standard deviations and mixing parameters as vector:

```
mu    <- c(-2, 5, 11)
sigma <- c(2.2, 1.4, 2.9)
prob  <- c(0.4, 0.25, 0.35)
n     <- 10000
```

  - Generate mixture by:

```
g <- sample(length(mu), n, replace=TRUE, p=prob)
x <- rnorm(n, mean = mu[g], sd = sigma[g])
hist(x, breaks = 25)
```



Ex. 5: Type I error of test under wrong distribution

- Given n independent and identically distributed observations $X_1, \ldots, X_n$ with mean μ, one can test H_0: μ=0 versus H_1: μ>0 with the one-sample t-test

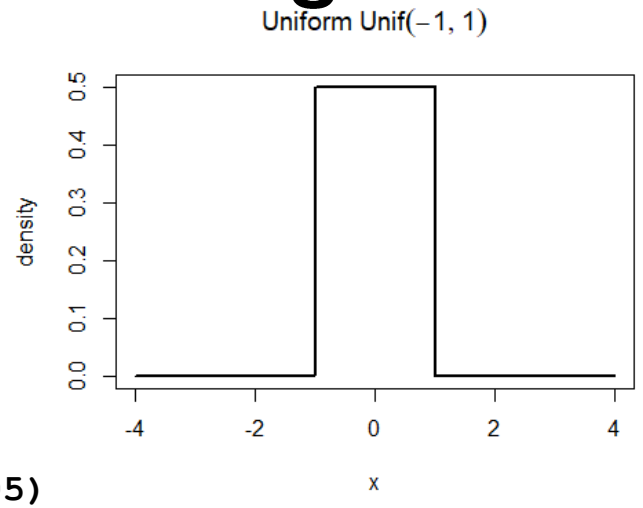
$$\text{reject } H_0 \text{ if and only if } \frac{\sqrt{n}\bar{x}}{s_x} > t_{n-1;1-\alpha}$$

- Assumption for test: normal distribution of observations
- How sensitive is t-test if observations not normal?
- We focus on H_0 first: Can type I error be larger than α (such that it matters) for certain distributions?
- Idea:
  - Choose some distributions with mean=0, simulate n repetitions, perform t-test, and record if rejected
  - Repeat this s times and check rejection rate

Ex. 5: Type I error of test under wrong distribution

- For n=10, simulate rejection rate for Unif[-1,1]

```
#Simulation of one sample t-test
s     <- 100000
n     <- 10
count <- 0
for (sim in 1:s)
{
  x      <- runif(n, min = -1, max = 1)
  reject <- (t.test(x, alternative = "greater")$p.value < 0.05)
  count  <- count + reject
}
#Rejection rate estimate:
rre   <- count/s
```



This is 1 if the condition in (...) is true, otherwise it is 0

- Note that there are possibilities to make simulation more efficient (e.g., by avoiding the loop) – see code on homepage

Ex. 5: Type I error of test under wrong distribution

```
s     <- 100000
n     <- 10
count <- 0
for (sim in 1:s)
{
  x      <- runif(n, min = -1, max = 1)
  reject <- (t.test(x, alternative = "greater")$p.value < 0.05)
  count  <- count + reject
}
rre   <- count/s
```

- Precision of result?

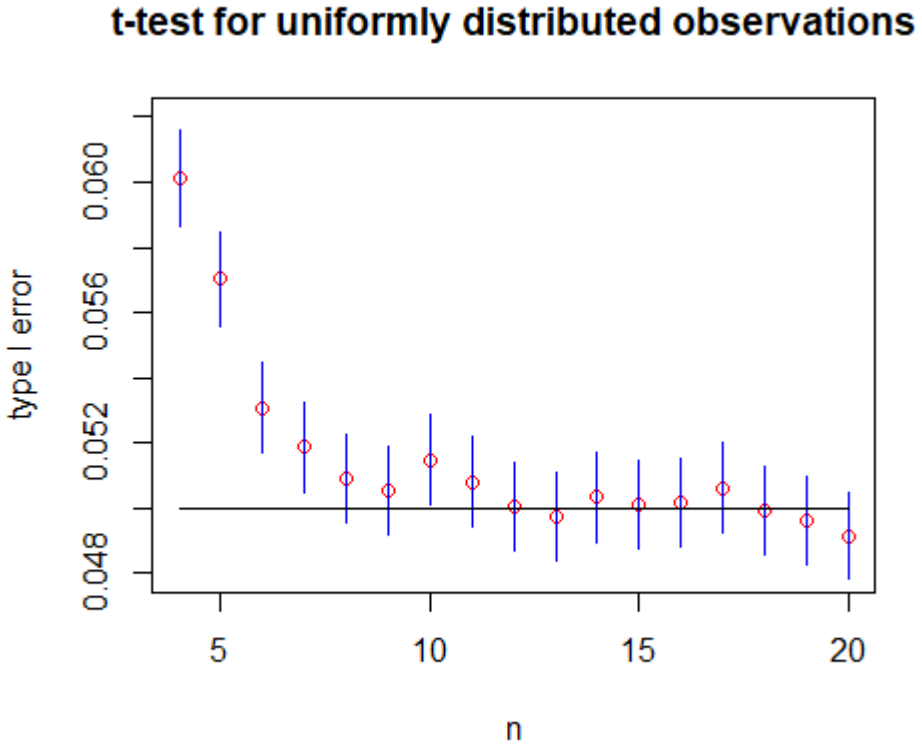
p = true rejection rate; reject$\sim Bin(1, p)$, count$\sim Bin(s = 100000, p)$

$$Var(\text{count}) = p(1-p)s, Var\left(\frac{\text{count}}{s}\right) = \frac{p(1-p)}{s}, sd(\text{rre}) = \sqrt{\frac{p(1-p)}{s}}$$

≈ 0.0007 for $p = 0.05$.

Ex. 5: Type I error of test under wrong distribution

- Simulated rejection rate for Unif[-1,1] for n = 4, 5, …, 20 with 95%-simulation-error-CIs based on 100 000 sim. for each n
- One more loop for n used
- Took ~1 min to simulate



Ex. 5: Type I error of test under wrong distribution

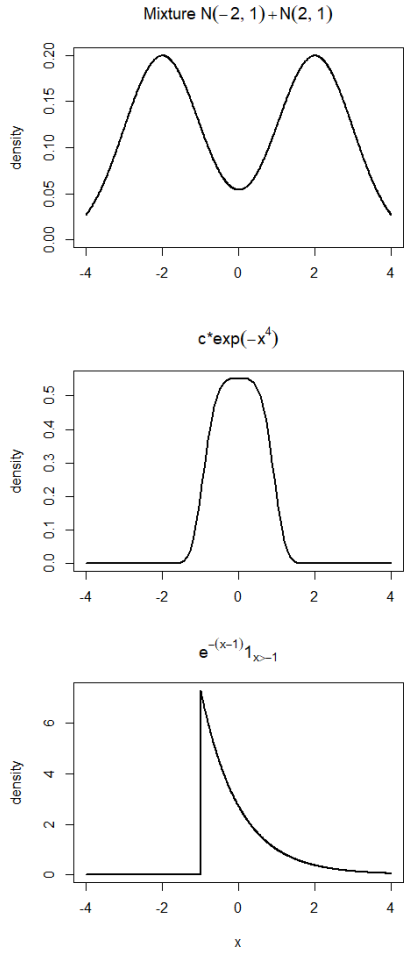
- Again, rejection rate for n=10, but for:

a) An equal mixture of N(-2,1) and N(2,1),

b) Distribution with density: $f(x)=c\ \exp(-x^4)$,

c) Distribution with density: $f(x)=e^{-(x-1)}\ 1_{\{x>-1\}}$

Which simulation method in each case?



Sampling importance resampling (SIR)

- Methods considered so far generate intended distribution **exactly**
- Sampling importance resampling (SIR) is **approximate** method (this approximation is often fully ok)

- Use again envelope-function g, but do not longer require the envelope being larger than f everywhere
- If desired to draw n observations following f, start with sampling m independent observations following g
  (recommendation: $m \geq 10n$)
- Resample then n from these m as described below

Sampling importance resampling (SIR)

1. Sample m (≥ 10n) random variables $Y_1, \ldots, Y_m$ from g
2. Calculate standardized importance weights

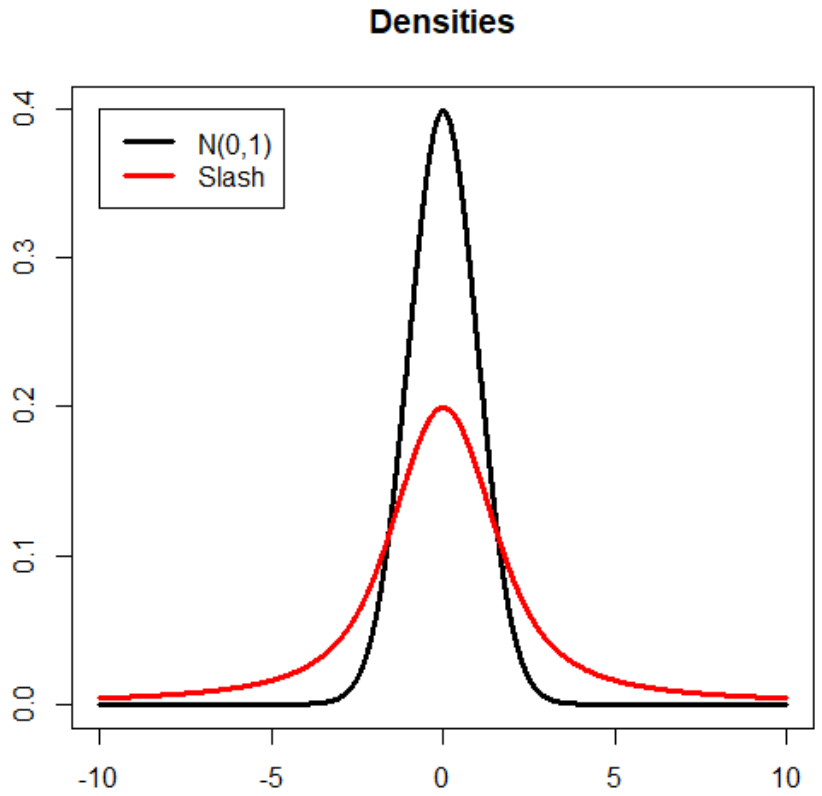
$$w(Y_i) = \frac{f(Y_i)/g(Y_i)}{\sum_{j=1}^{m} f(Y_j)/g(Y_j)}$$

   for all m random draws Y_i from g.
3. Resample $X_1, \ldots, X_n$ from $Y_1, \ldots, Y_m$ with replacement with probabilities $w(Y_1), \ldots, w(Y_m)$

- $X_1, \ldots, X_n$ follow then approximately f
- Note: f need to be known only up to a constant (constant cancels out in calculation of standardized weights)

The slash distribution

- A random variable Y has slash distribution if Y = X/U with X ~ N(0,1) and U ~ Unif(0,1) independently

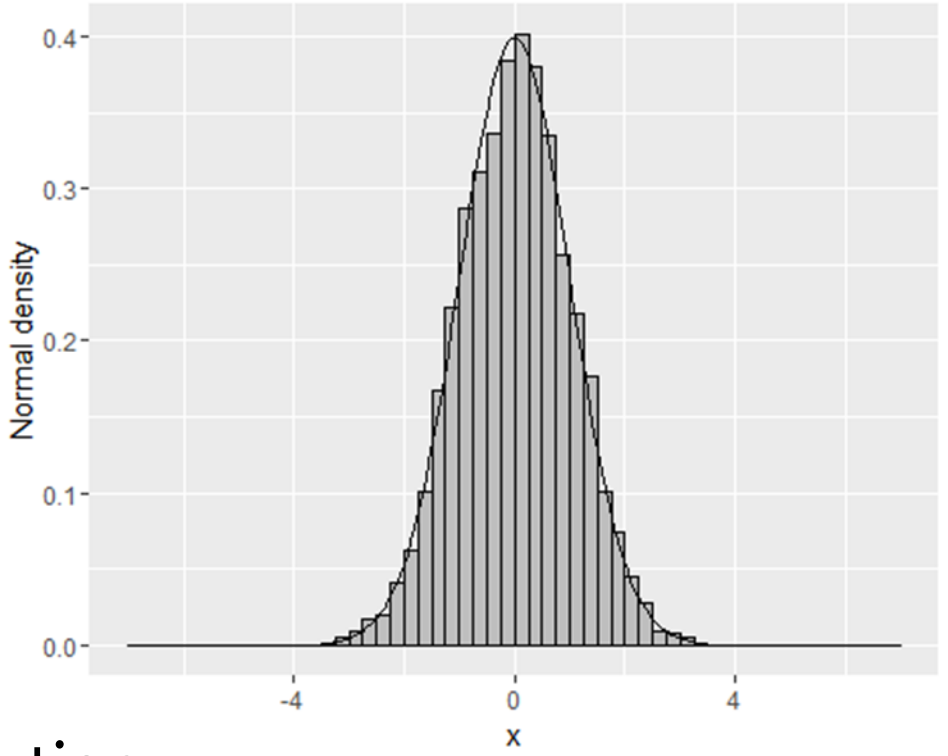


Sampling importance resampling (SIR) – Illustration

(Thanks to Yuliya Leontyeva for code and illustration on this slide!)

- Example 6: Use slash distribution as SIR envelope g to generate random variables following standard normal density f

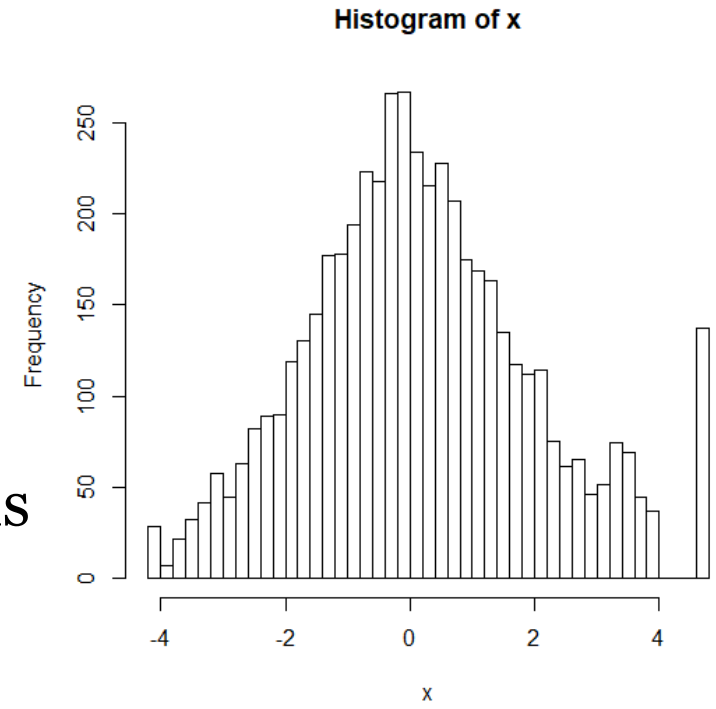
```
library("extraDistr") # used for computation of slash
                      # density and simulation

sir <- function(m, n)
{
  # m - sample size from the envelope distribution
  # n - resample size
  # relative to n, m should be large
  y <- rslash(m) # sample candidates Y1,...Ym iid from g
  w <- dnorm(y)/dslash(y)
  w <- w/sum(w) # calculate the standardized weights
  x <- sample(y, n, replace=TRUE, prob=w) # resample with
                                          # probabilities = w

  return(x)
}
x <- sir(100000, 5000)
```



The simulated data follows well a normal distribution

Sampling importance resampling (SIR) – illustration

- Method worked well since envelope (slash distribution) had heavier tails than target distribution (standard normal)
- If we run SIR to generate the slash distribution with standard normal as envelope, no observations are generated at tails
- Lowest and highest values in Y-sample receive high weights (overrepresentation in X-sample)
- **Recommendation:** Use envelopes with heavier tails (or equally heavy) than the target distribution



Markov chain Monte Carlo (MCMC), see GH 7.1, 7.3

- The algorithms considered so far generate sequences of **independent** observations which follow the target distribution exactly or approximately (sampling importance resampling)
- We will now consider a method which generates a sequence of **dependent** observations which follow the target distribution approximately

- The next observation (t+1) will be generated based on a proposal distribution g which depends on the current observation (t), i.e. $g(\cdot|X^{(t)})$
- Since $X^{(t+1)}$ depends on $X^{(t)}$ but not on earlier observations, the sequence $(X^{(t)})$ is a Markov chain

MCMC – Metropolis-Hastings algorithm

- A general method to generate the Markov chain is the Metropolis-Hastings (MH) algorithm
- A starting value $x^{(0)}$ is generated from some starting distribution
- Given observation $x^{(t)}$, generate $x^{(t+1)}$ as follows:

1. Sample a candidate x^* from a proposal distribution $g(\cdot|x^{(t)})$
2. Compute the MH ratio $R(x^{(t)}, x^*) = \frac{f(x^*)\, g(x^{(t)}|x^*)}{f(x^{(t)})\, g(x^*|x^{(t)})}$
3. Sample $x^{(t+1)}$ according to

$$x^{(t+1)} = \begin{cases} x^*, \text{with probability } \min\{R(x^{(t)}, x^*), 1\} \\ x^{(t)}, \text{otherwise} \end{cases}$$

4. If more observations needed, set `t <- t+1`; go to 1

Metropolis algorithm
Special case when g is symmetric:
$g(x^*|x^{(t)}) = g(x^{(t)}|x^*)$

$= \frac{f(x^*)}{f(x^{(t)})}$

Simulated annealing

- Start value $x^{(0)}$; Stage j=0,1,2,... has m_j iterations; set j=0
- Given iteration $x^{(t)}$, generate $x^{(t+1)}$ as follows:

1. Sample a candidate x^* from a proposal distribution $p(\cdot|x^{(t)})$
2. Compute $h(x^{(t)}, x^*) = \exp(\frac{g(x^*)-g(x^{(t)})}{\tau_j})$
3. Define next iteration $x^{(t+1)}$ according to

$$x^{(t+1)} = \begin{cases} x^*, \text{with probability } \min\{h(x^{(t)}, x^*), 1\} \\ x^{(t)}, \text{otherwise} \end{cases}$$

4. Set `t <- t+1` and repeat 1.-3. m_j times
5. Update $\tau_j = \alpha(\tau_{j-1})$ and $m_j = \beta(m_{j-1})$; set `j <- j+1`; go to 1

$g(x^{(t)}) - g(x^*)$ for minimisation

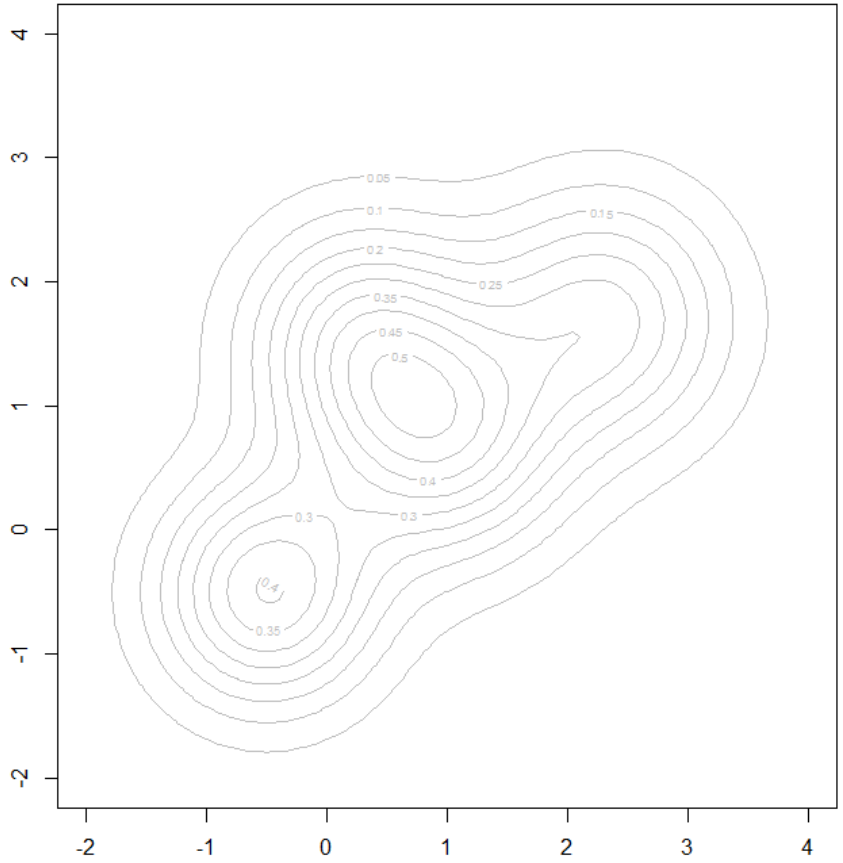
τ_j is temperature; function α should slowly decrease it; function β should be increasing

Simulated annealing and Metropolis algorithm

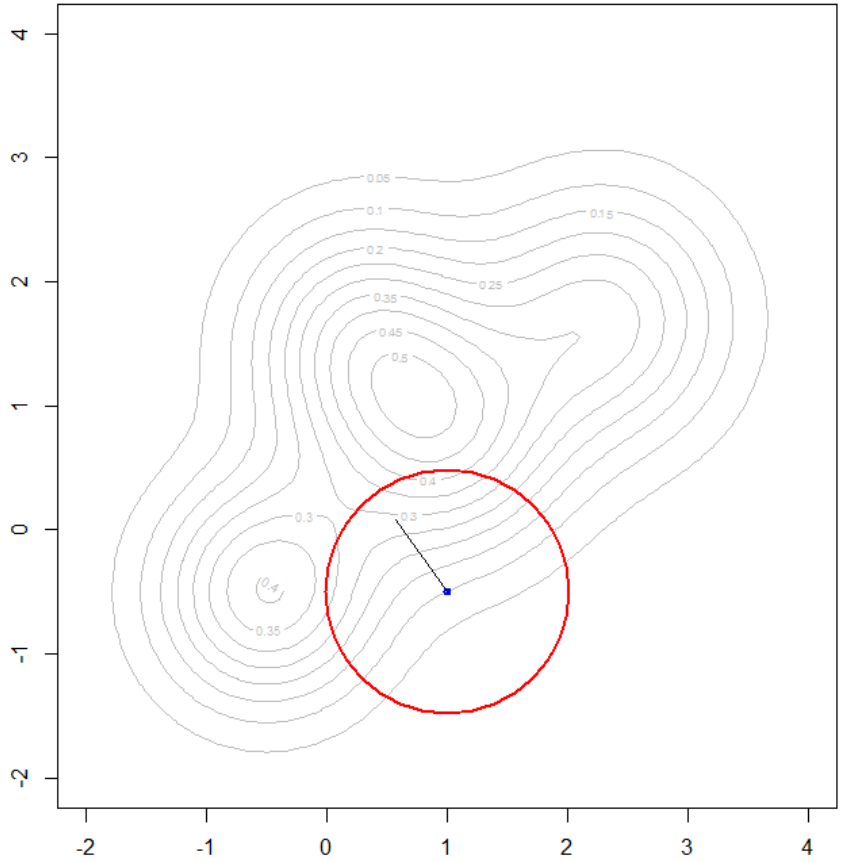
- For fixed temperature τ, simulated annealing algorithm is a Metropolis algorithm
- Kirkpatrick et al. (1983) proposed name simulated annealing for using it as optimisation method
- $h\left(x^{(t)}, x^*\right) = \exp\left(\frac{g\left(x^{(t)}\right) - g(x^*)}{\tau_j}\right) = \frac{\exp\left(-\frac{g(x^*)}{\tau_j}\right)}{\exp\left(-\frac{g\left(x^{(t)}\right)}{\tau_j}\right)} = \frac{f(x^*)}{f\left(x^{(t)}\right)} = R(x^{(t)}, x^*)$
- Key ingredient of Metropolis and simulated annealing alg.: Markov chain $\boldsymbol{x^{(t)}}$ **has limiting stationary distribution** $\boldsymbol{f}$; for a proof see e.g. Koski (2009)
- Requirement for all: $x^{(t)}$ irreducible and aperiodic chain

Metropolis alg. – Ex.7

- For illustration, we consider two-dimensional distribution with density f according to contour lines in figure (extended **example from L3**)
- Proposal distribution
  $g(x^*|x^{(t)}) = g(x^{(t)}|x^*)$
  $= \frac{1}{\pi r^2}\mathbf{1}\{\|x^{(t)} - x^*\| < r\}$
  for some constant r (here=1)

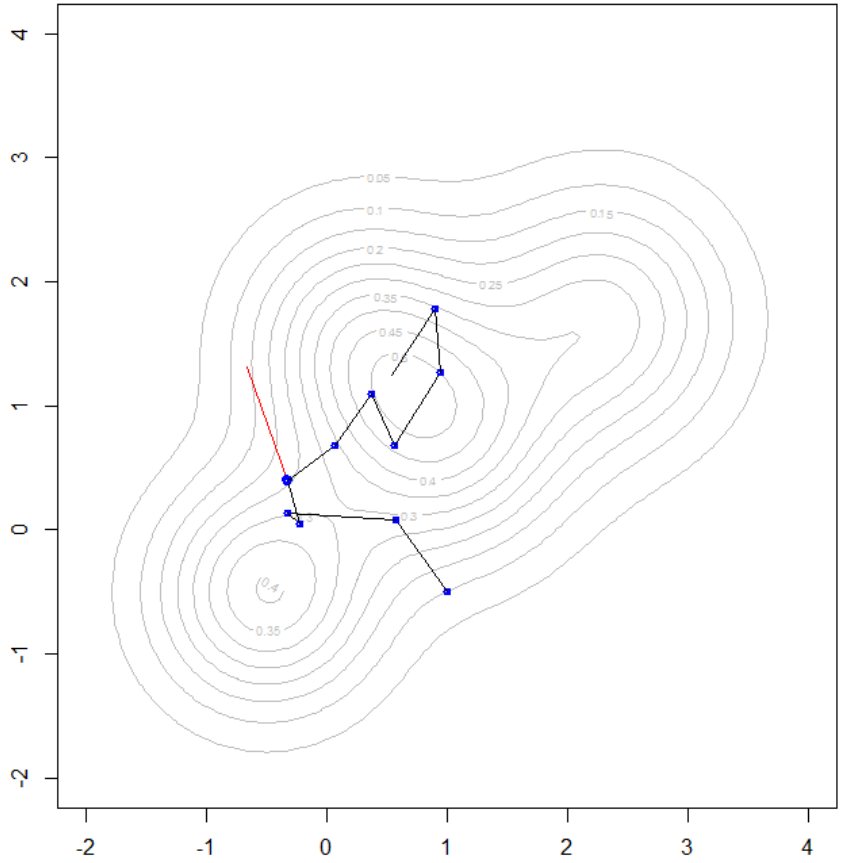


Metropolis alg. – Ex.7



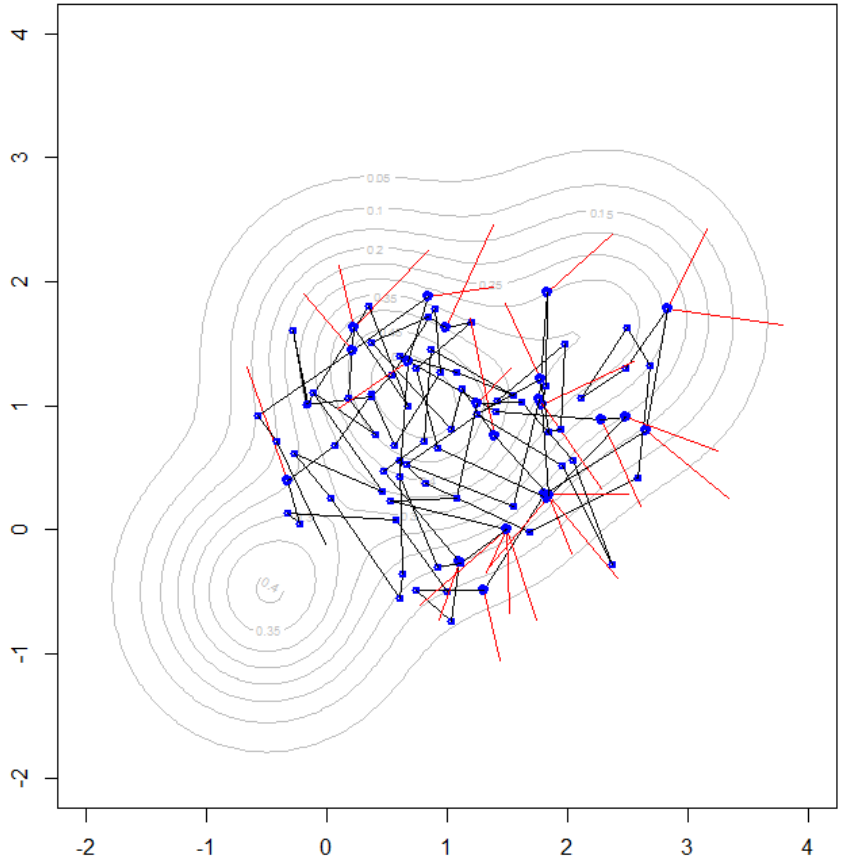
- Proposal distribution
  $g(x^*|x^{(t)}) = g(x^{(t)}|x^*)$
  $= \frac{1}{\pi r^2}\mathbf{1}\{\|x^{(t)} - x^*\| < r\}$
  for some constant r (here=1)
- Start here with $x^{(0)}=(1,-0.5)$
- Randomize uniformly on unit circle around $x^{(0)}$ (proposal distribution); result $x^*=(0.58,0.08)$
- $f(x^*)=0.296 > f(x^{(0)}) = 0.098$; so this was an uphill step and is automatically accepted ($R\left(x^{(t)}, x^*\right) = \frac{f(x^*)}{f\left(x^{(t)}\right)} > 1$)

Metropolis alg. – Ex.7



- $x^{(0)}=(1,-0.5)$
- Uphill steps: $x^{(1)}=(0.58,0.08)$
- $x^{(2)}=(-0.33,0.13)$
- $x^{(3)}=(-0.23,0.05)$
- Then downhill step proposed: $x^{*}=(-0.32,0.4)$, $R\left(x^{(t)},x^{*}\right)=\frac{f(x^{*})}{f\left(x^{(t)}\right)}=0.774$
- Random Unif(0,1) generated: 0.573 and since this is smaller than R=0.774, $x^{(4)}=x^{*}=(-0.32,0.4)$ is accepted
- Again downhill step proposed: $x^{*}=(-0.67,1.31)$, $R\left(x^{(t)},x^{*}\right)=\frac{f(x^{*})}{f\left(x^{(t)}\right)}=0.560$; random Unif(0,1): 0.890 and rejection of x^{*}
- $x^{(5)}=x^{(4)}=(-0.32,0.4)$

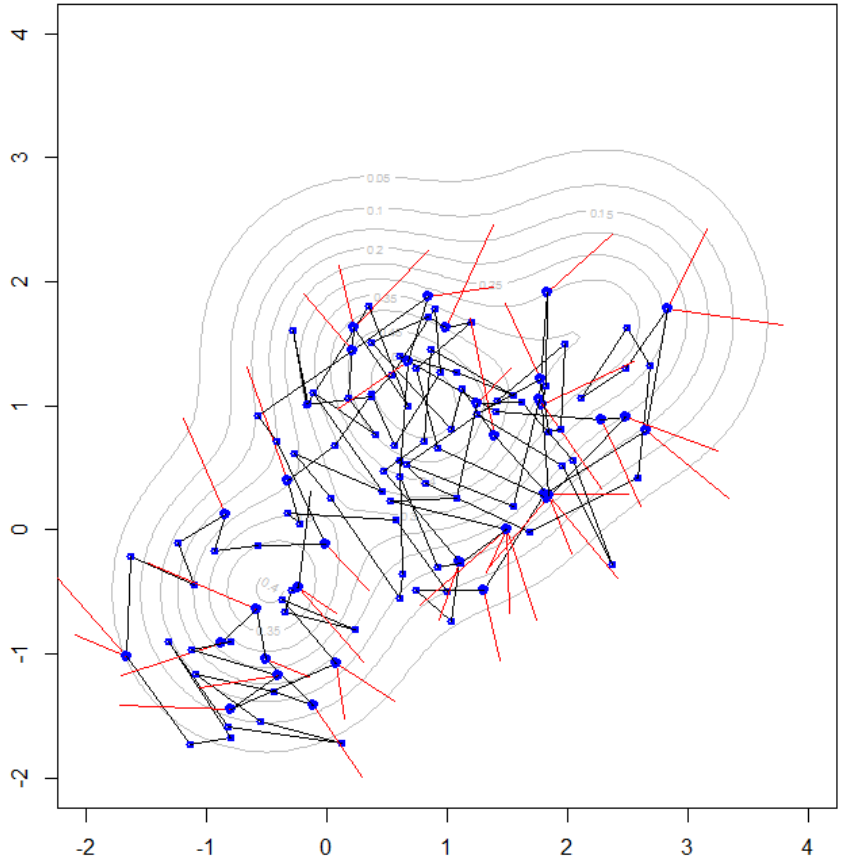
Metropolis alg. – Ex.7



- After several additional iterations (see red lines for rejected proposals), one part of the distribution was explored to a good extend
- Since uphill steps preferred, part of distribution with local maximum at (-0.5,-0.5) is not yet ”detected” at all
- Occasionally, the path will arrive at this part as well

Metropolis alg. – Ex.7

- Now, larger parts of distribution explored



- A couple of animations can be found on: https://chi-feng.github.io/mcmc-demo/app.html#RandomWalkMH,standard (choose Algorithm: RandomWalkMH)

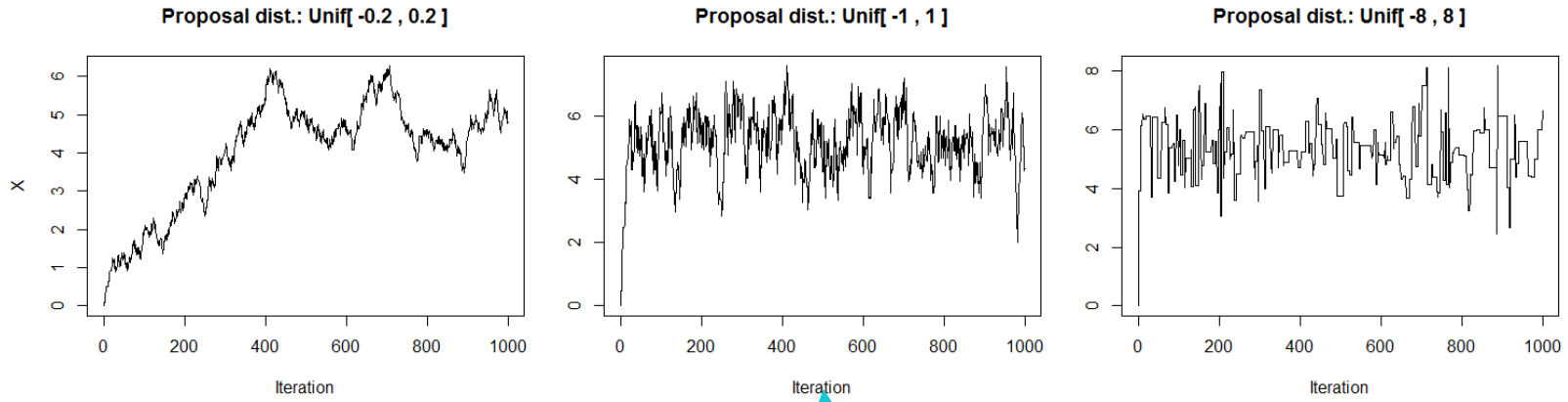
Metropolis algorithm - Example 8

(compare Givens and Hoeting, ex. 5.3)

- Consider Bayesian estimation of μ based on $N(\mu,3^2/7)$ likelihood for μ and Cauchy(5,2) prior; observed mean=5.38
- The posterior density is proportional to product of likelihood and prior density
- **Use MCMC to generate random samples following the posterior density**
- Based on these random samples, one can e.g.
  - determine posterior probability that $2 \leq \mu \leq 8$
  - determine mean and variance of posterior

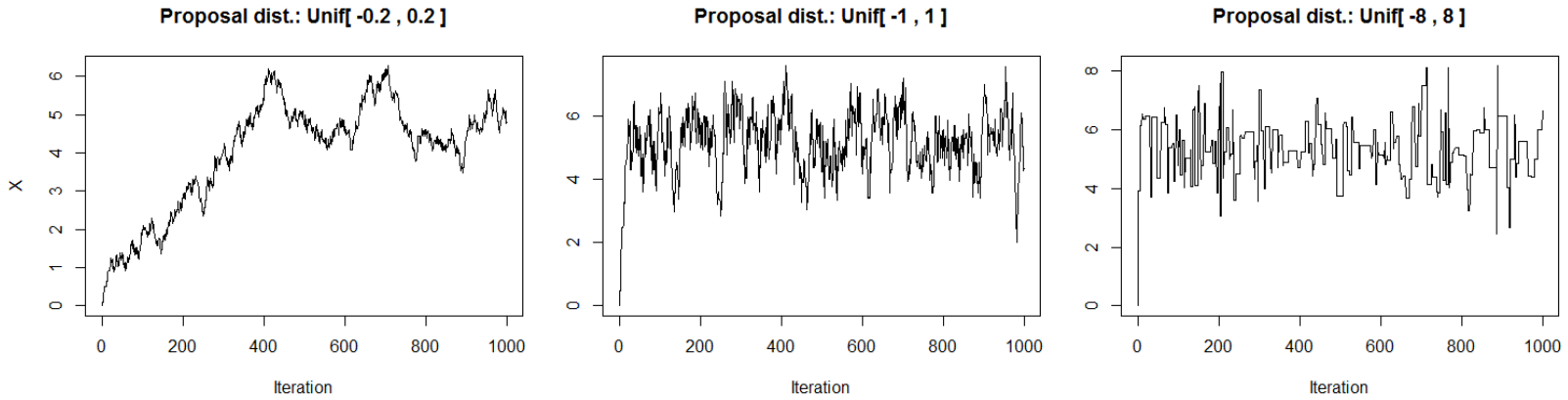
Metropolis algorithm - Example 8

- We use starting value $x^{(0)}=0$, s=1000 iterations and following proposal distributions $g(\cdot|x^{(t)})$:
  $x^{(t)}$+Unif[-0.2,0.2], $x^{(t)}$+Unif[-1,1], $x^{(t)}$+Unif[-8,8]
- **Sample path plots** show simulated values $x^{(t)}$ vs. iteration number t



Best choice

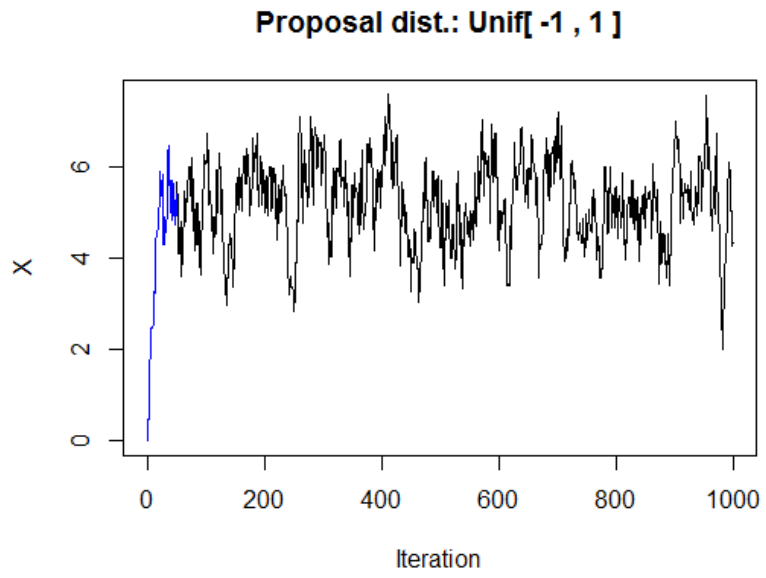
Metropolis algorithm - Example 8

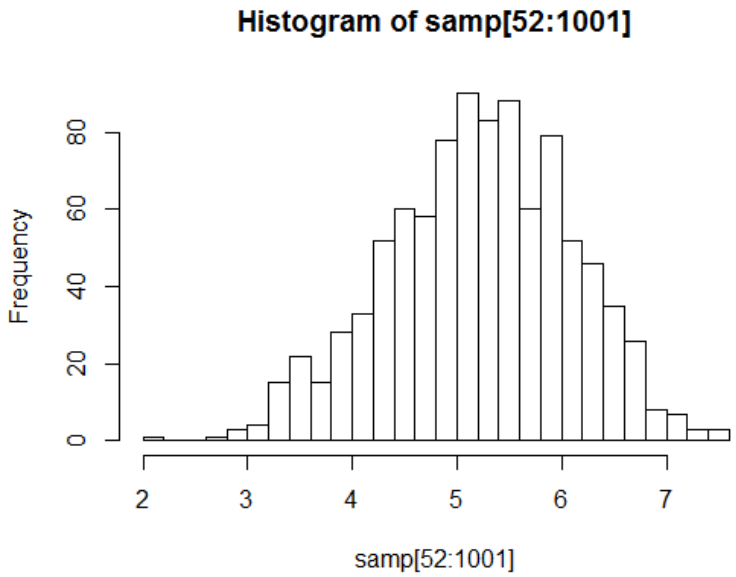


- Count “acceptance rate” (=proportion accepted proposals)
- Here: 98% 78% 18%
- Best results for 44% (uni-dim. case) to 23.4% (high dim. case) acceptance probability (theory based on normal target and proposal functions, see Givens and Hoeting, Chapter 7.3, for references about that)
- For multimodal functions lower acceptance probabilities might be good

Metropolis algorithm - Example 8

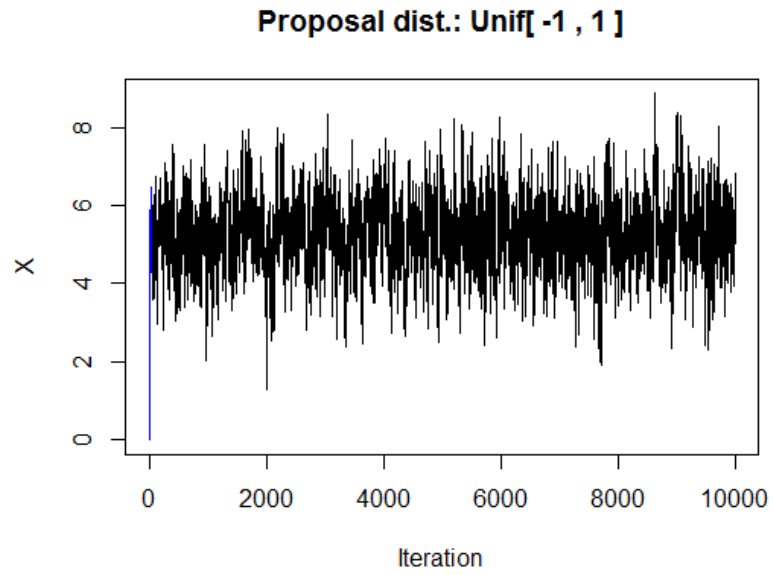
- Based on sample path plots, we might choose Unif[-1,1] as proposal distribution
- Often, one wants to discard initial samples (**burn-in** period) which highly depend on starting value, e.g. 50 values + $x^{(0)}$

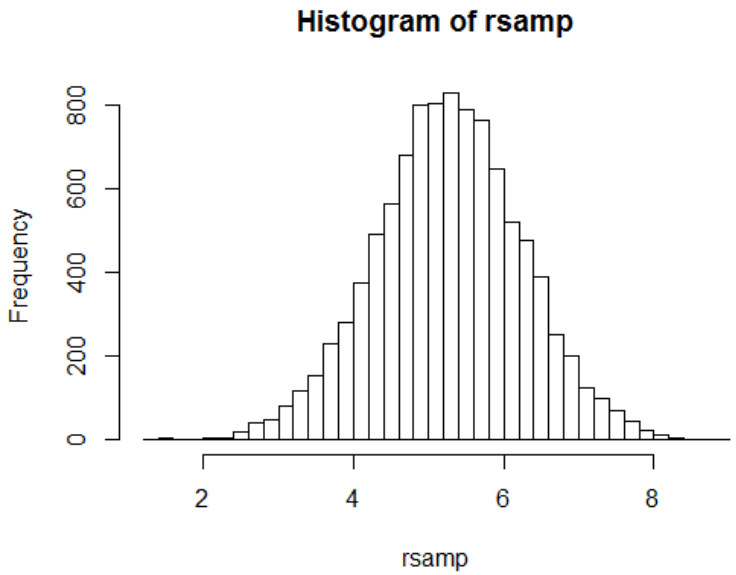




Metropolis algorithm - Example 8

- For s=10 000 iterations and burn-in of 50, we obtain

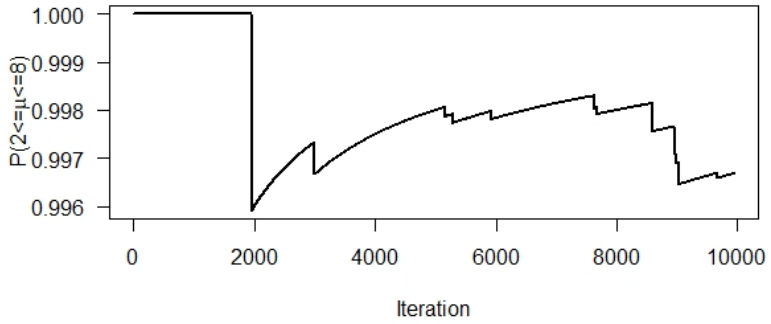




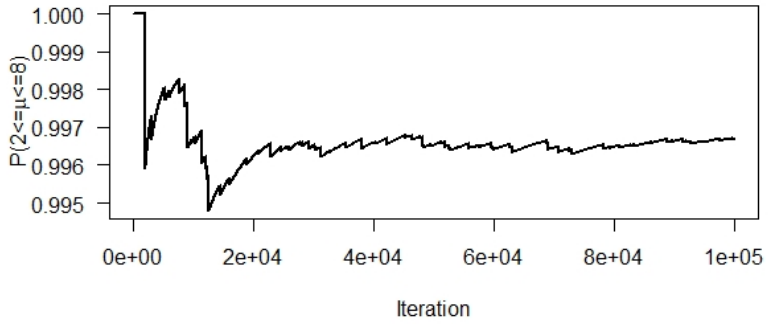
- Monte Carlo estimate for $P(2 \leq \mu \leq 8)$ is 0.9967
  (Monte Carlo standard error=$\sqrt{0.9967 * 0.0033/9950} = 0.0006$)
- Estimated mean = 5.26, standarddeviation =0.99

Metropolis algorithm - Example 8

- Were s=10 000 iterations enough to ensure convergence?
- Can depend on the purpose …
- E.g. for estimating P(2≤μ≤8)
- One can monitor cusum/convergence plots showing estimate versus iterations (see Givens and Hoeting, ch.7.3.1.1)
- After 10 000 iterations



After 100 000 iterations



- After 10 000 iterations, we might not be happy with the left graph; we run longer and are happy with 100 000

Metropolis-Hastings with independent proposals

- Other proposal distributions g possible (not necessarily symmetric), e.g. independent proposals
- Proposal distribution depends not on previous value, $g(\cdot|\mathrm{x}^{(t)}) = g(\cdot)$
- The MH ratio is $R\left(x^{(t)}, x^*\right) = \frac{f(x^*)\, g\left(x^{(t)}|x^*\right)}{f\left(x^{(t)}\right) g\left(x^*|x^{(t)}\right)} = \frac{f(x^*)/g(x^*)}{f\left(x^{(t)}\right)/g(x^{(t)})}$

- A possible application is for Bayesian analysis (f is the posterior) with proposal distribution g being the prior distribution
- f/g is then the likelihood

Markov chain Monte Carlo

- In Givens and Hoeting (2013), Chapter 7 and 8, more about Markov chain Monte Carlo algorithms